

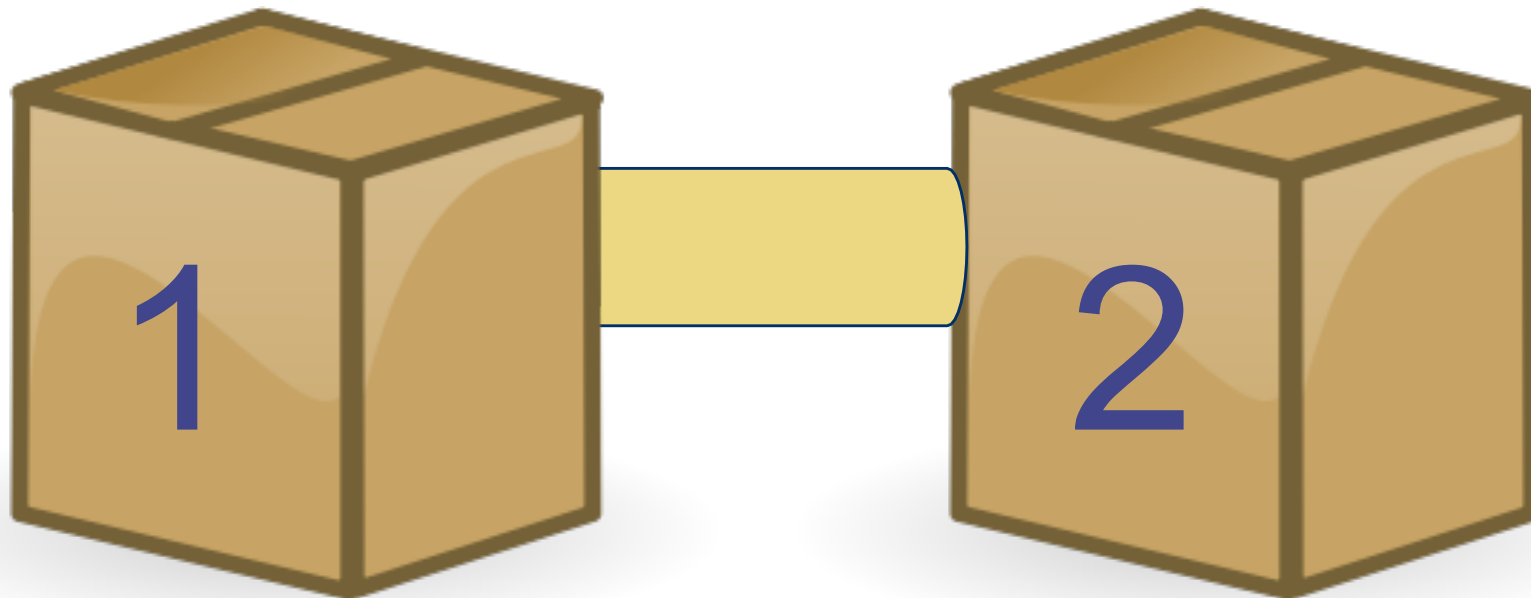
COMP9334

Capacity Planning for Computer Systems and Networks

Week 3: Queues with Poisson arrivals

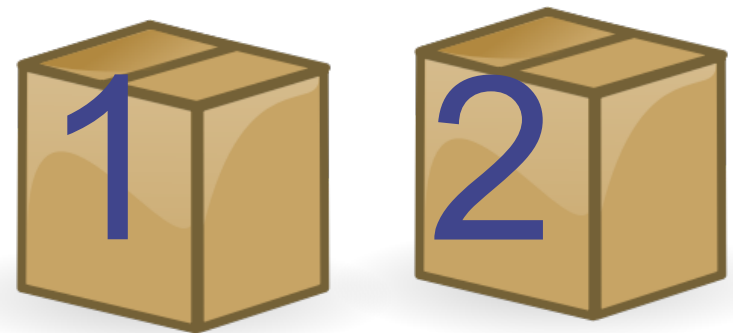
Pre-lecture exercise: Where is Felix? (1)

- You have two boxes: Box 1 and Box 2, as well as a cat called Felix
- The two boxes are connected by a tunnel
- Felix likes to hide inside these boxes and travels between them using the tunnel.
- Felix is a very fast cat so the probability of finding him in the tunnel is zero
- You know Felix is in one of the boxes but you don't know which one



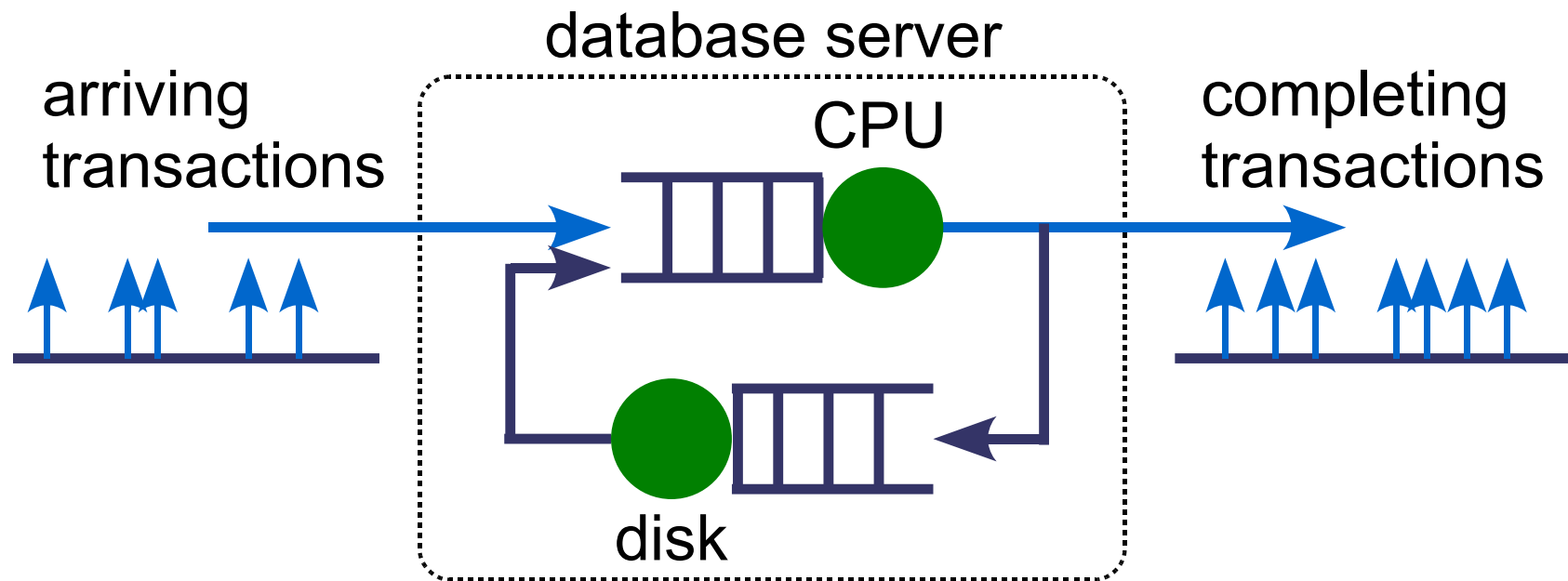
Pre-lecture exercise: Where is Felix? (2)

- Notation:
 - $\text{Prob}[A]$ = probability that event A occurs
 - $\text{Prob}[A \mid B]$ = probability that event A occurs given event B
- You do know
 - Felix is in one of the boxes at times 0 and 1
 - $\text{Prob}[\text{Felix is in Box 1 at time 0}] = 0.3$
 - $\text{Prob}[\text{Felix will be in Box 2 at time 1} \mid \text{Felix is in Box 1 at time 0}] = 0.4$
 - $\text{Prob}[\text{Felix will be in Box 1 at time 1} \mid \text{Felix is in Box 2 at time 0}] = 0.2$
- Calculate
 - $\text{Prob}[\text{Felix is in Box 1 at time 1}]$
 - $\text{Prob}[\text{Felix is in Box 2 at time 1}]$



Week 1:

- Modelling a computer system as a network of queues
- Example: Open queueing network consisting of two queues

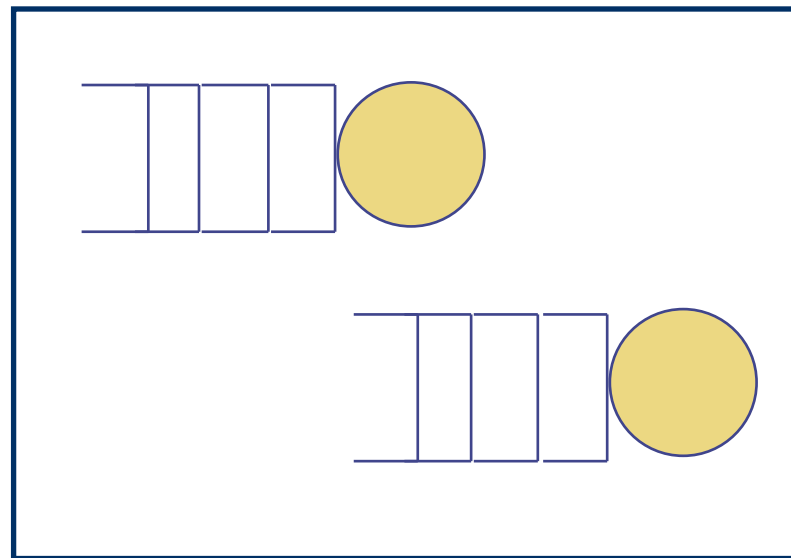


Week 2:

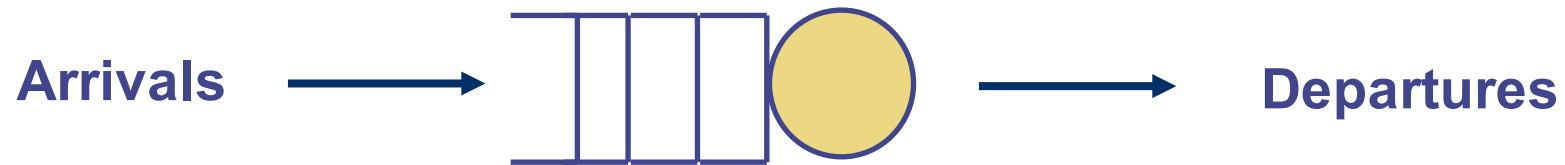
- Operational analysis
 - Measure #completed jobs, busy time etc
 - Operational quantities: utilisation, response time, throughput etc.
 - Operational laws relate the operational quantities
- Bottleneck analysis

Little's Law

- Applicable to any “box” that contains some queues or servers
- Mean number of jobs in the “box” =
Mean response time x Throughput
- We will use Little's Law in this lecture to derive the mean response time
 - We first compute the mean number of jobs in the “box” and throughput

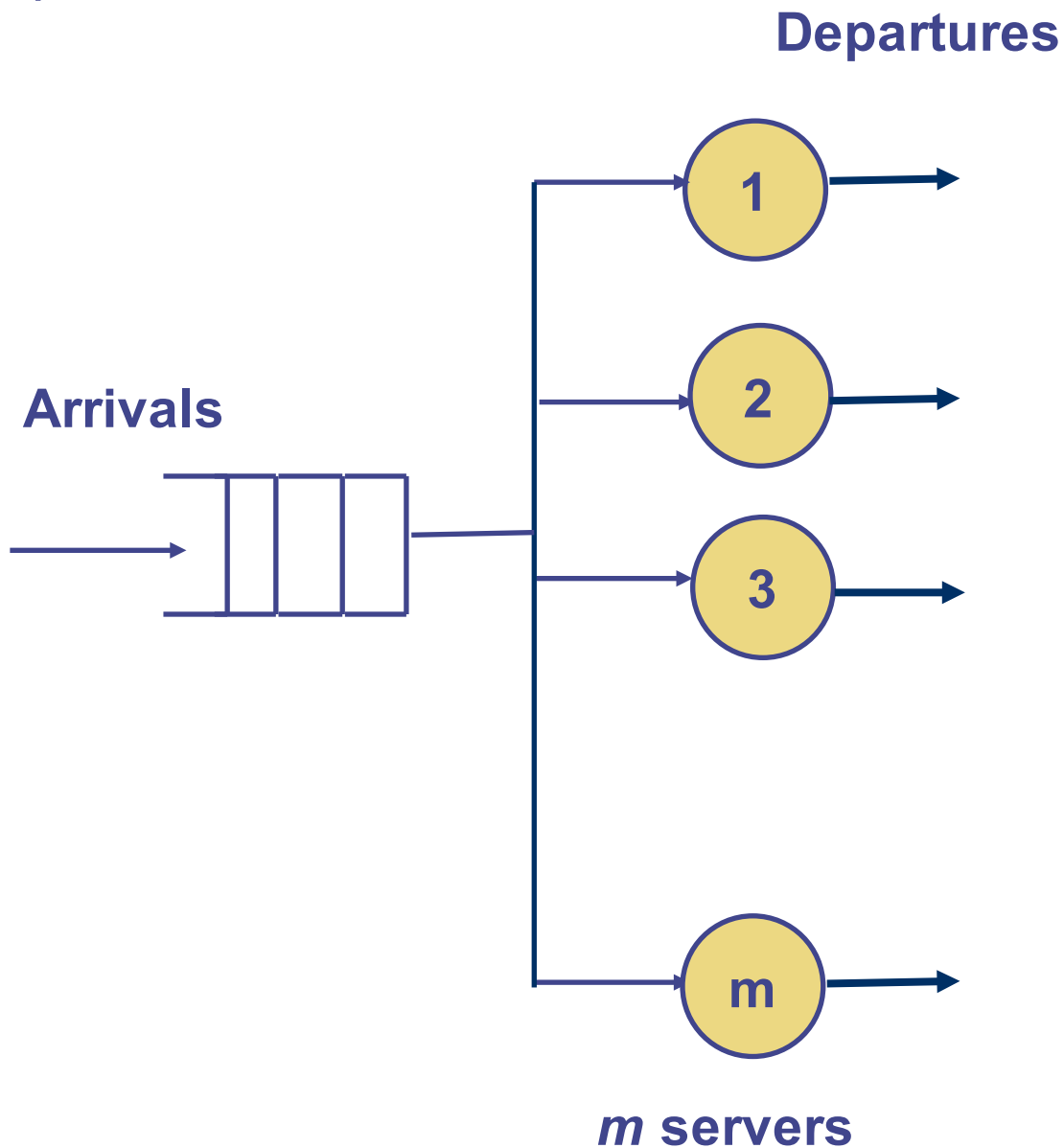


This week (1)



- Open, single server queues and
- How to find:
 - Waiting time
 - Response time
 - Mean queue length etc.
- The technique to find waiting time etc. is called *Queueing Theory*

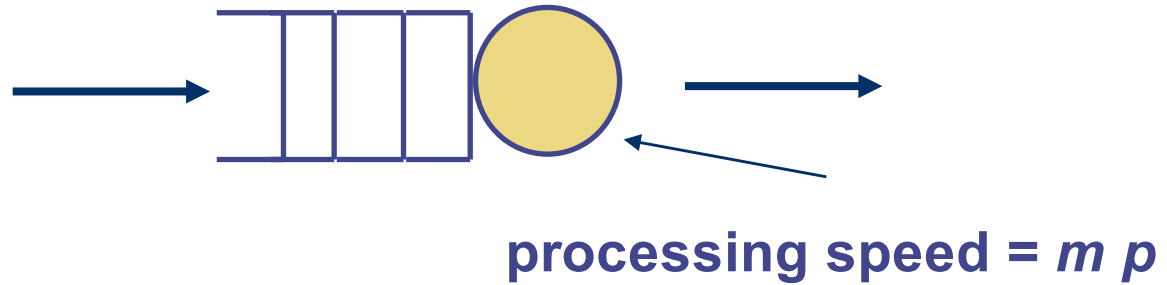
This week (2)



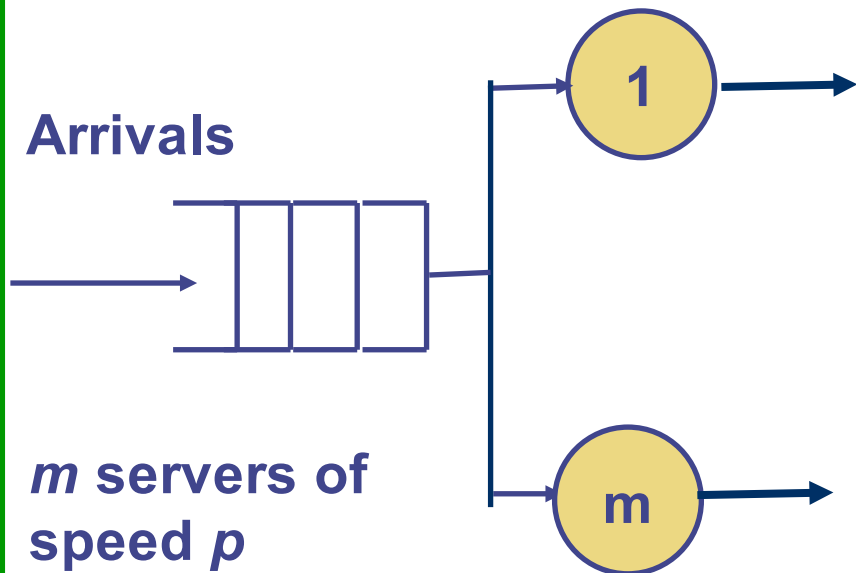
- Open, multi-server queue
- How to find:
 - Waiting time
 - Response time
 - Mean queue length etc.

What will you be able to do with the results?

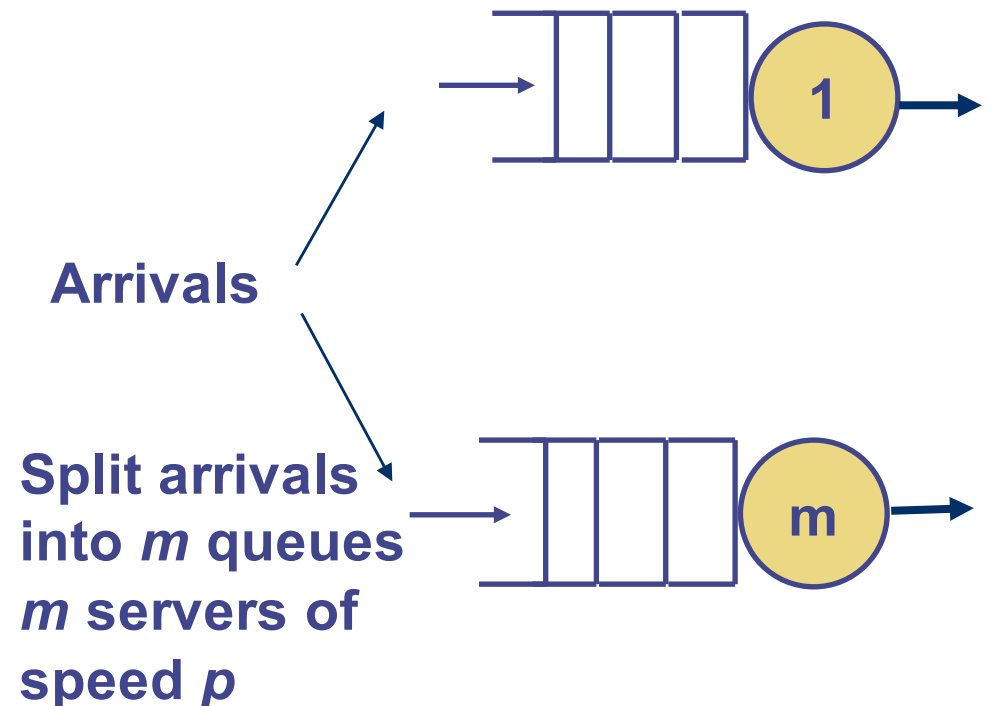
Configuration 1:



Configuration 2:



Configuration 3:

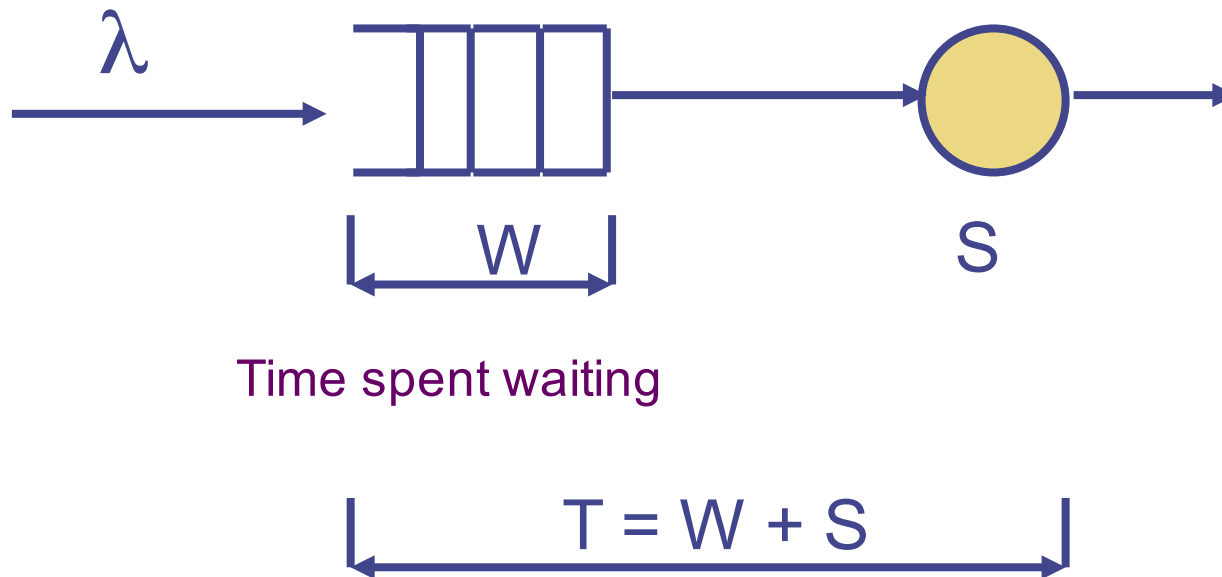


Which configuration has the best response time?

Be patient

- We will show how we can obtain the response time
 - It takes a number of steps to obtain the answer
- It takes time to stand in a queue, it also takes time to derive results in queuing theory!

Single Server Queue: Terminology



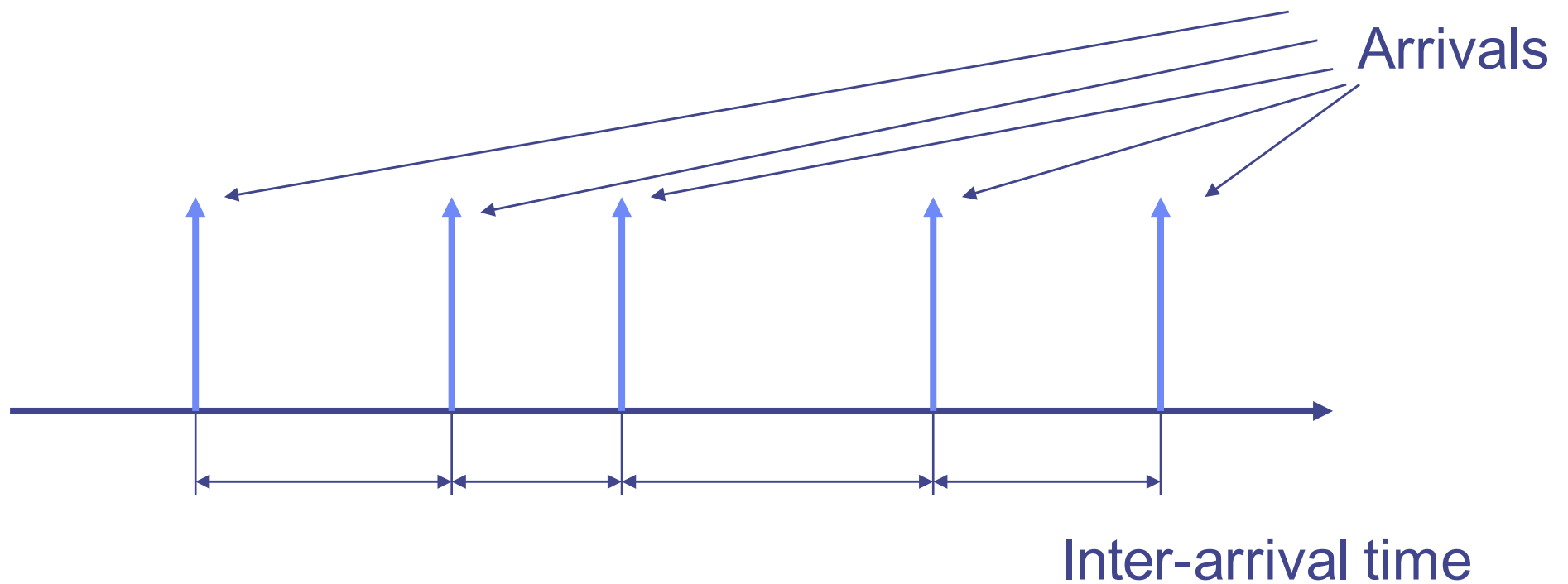
Response Time T
= Waiting time W + Service time S

Note: We use T for response time because this is the notation in many queueing theory books. For a similar reason, we will use ρ for utilisation rather than U .

Single server system

- In order to determine the response time, you need to know
 - The inter-arrival time probability distribution
 - The service time probability distribution
- Possible distributions
 - Deterministic
 - Constant inter-arrival time
 - Constant service time
 - Exponential distribution
- We will focus on exponential distribution

Exponential inter-arrival with rate λ



We assume that successive arrivals are independent

Probability that inter-arrival time is between x and $x + \delta x$
 $= \lambda \exp(-\lambda x) \delta x$

Poisson distribution (1)

- The following are equivalent
 - The inter-arrival time is independent and exponentially distributed with parameter λ
 - The number of arrivals in an interval T is a Poisson distribution with parameter λ

$$Pr[k \text{ arrivals in a time interval } T] = \frac{(\lambda T)^k \exp(-\lambda T)}{k!}$$

- Mean inter-arrival time = $1 / \lambda$
- Mean number of arrivals in time interval $T = \lambda T$
- Mean arrival rate = λ

Poisson distribution (2)

- Poisson distribution arises from a large number of independent sources
 - An example from Week 2:
 - N customers, each with a probability of p per unit time to make a request.
 - This creates a Poisson arrival with $\lambda = Np$
- Another interpretation of Poisson arrival:
 - Consider a small time interval δ
 - This means δ^n (for $n \geq 2$) is negligible
 - Probability [no arrival in δ] = $1 - \lambda \delta$
 - Probability [1 arrival in δ] = $\lambda \delta$
 - Probability [2 or more arrivals in δ] ≈ 0
- This interpretation can be derived from:

$$Pr[k \text{ arrivals in a time interval } T] = \frac{(\lambda T)^k \exp(-\lambda T)}{k!}$$

Service time distribution

- Service time = the amount of processing time a job requires from the server
- We assume that the service time distribution is exponential with parameter μ
 - The probability that the service time is between t and $t + \delta t$ is:

$$\mu \exp(-\mu t) \delta t$$

- Here: μ = service rate = $1 / \text{mean service time}$
- Another interpretation of exponential service time:
 - Consider a small time interval δ
 - Probability [a job will finish its service in next δ seconds] = $\mu \delta$

Sample queueing problems

- Consider a call centre
 - Calls are arriving according to Poisson distribution with rate λ
 - The length of each call is exponentially distributed with parameter μ
 - Mean length of a call is $1/\mu$ (in, e.g. seconds)

Call centre:

Arrivals



m operators

If all operators are busy, the centre can put at most n additional calls on hold.

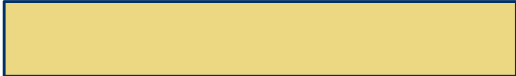
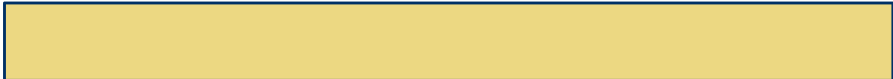
If a call arrives when all operators and holding slots are used, the call is rejected.

- Queueing theory will be able to answer these questions:
 - What is the mean waiting time for a call?
 - What is the probability that a call is rejected?

Road map

- We will start by looking at a call centre with one operator and no holding slot
 - This may sound unrealistic but we want to show how we can solve a typical queueing network problem
 - After that we go into queues that are more complicated

Call centre with 1 operator and no holding slots

- Let us see how we can solve the queuing problem for a very simple call centre with 1 operator and no holding slots
- What happens to a call that arrives when the operator is busy?
 - 
- What happens to a call that arrives when the operator is idle?
 - 
- We are interested to find the probability that an arriving call is rejected.

Arrivals



Call centre:

1 operator. No holding slot.

Solution (1)

- There are two possibilities for the operator:
 - Busy or
 - Idle
- Let
 - State 0 = Operator is idle (i.e. #calls in the call centre = ?)
 - State 1 = Operator is busy (i.e. #calls in the call centre = ?)

$P_0(t)$ = Prob. 0 call in the call centre at time t

$P_1(t)$ = Prob. 1 call in the call centre at time t

Solution (2)

We try to express $P_0(t + \Delta t)$ in terms of $P_0(t)$ and $P_1(t)$

- No call at call centre at $t + \Delta t$ can be caused by

- 
- 

**Question: Why do we NOT have to consider the following possibility:
No customer at time t & 1 customer arrives in $[t, t + \Delta t]$ &
the call finishes within $[t, t + \Delta t]$.**

Solution (3)

- Similarly, we can show that

$$P_1(t + \Delta t) = P_0(t)\lambda\Delta t + P_1(t)(1 - \mu\Delta t)$$

- If we let $\Delta t \rightarrow 0$, we have

$$\frac{dP_0(t)}{dt} = -P_0(t)\lambda + P_1(t)\mu$$

$$\frac{dP_1(t)}{dt} = P_0(t)\lambda - P_1(t)\mu$$

Solution (4)

- We can solve these equations to get

$$P_0(t) = \frac{\mu}{\lambda + \mu} - \frac{\mu}{\lambda + \mu} e^{-(\mu + \lambda)t}$$

$$P_1(t) = \frac{\lambda}{\lambda + \mu} + \frac{\mu}{\lambda + \mu} e^{-(\mu + \lambda)t}$$

- This is too complicated, let us look at **steady state** solution

$$P_0 = P_0(\infty) = \frac{\mu}{\lambda + \mu}$$

$$P_1 = P_1(\infty) = \frac{\lambda}{\lambda + \mu}$$

Solution (5)

- From the steady state solution, we have
 - The probability that an arriving call is rejected
 - = The probability that the operator is busy

- =
$$P_1 = \frac{\lambda}{\lambda + \mu}$$

- Let us check whether it makes sense
 - For a constant μ , if the arrival rate rate λ increases, will the probability that the operator is busy go up or down?
 - Does the formula give the same prediction?

An alternative interpretation

- We have derived the following equation:

$$P_0(t + \Delta t) = P_0(t)(1 - \lambda\Delta t) + P_1(t)\mu\Delta t$$

- Which can be rewritten as:

$$P_0(t + \Delta t) - P_0(t) = -P_0(t)\lambda\Delta t + P_1(t)\mu\Delta t$$

- At steady state:

Change in Prob in State 0 = 0

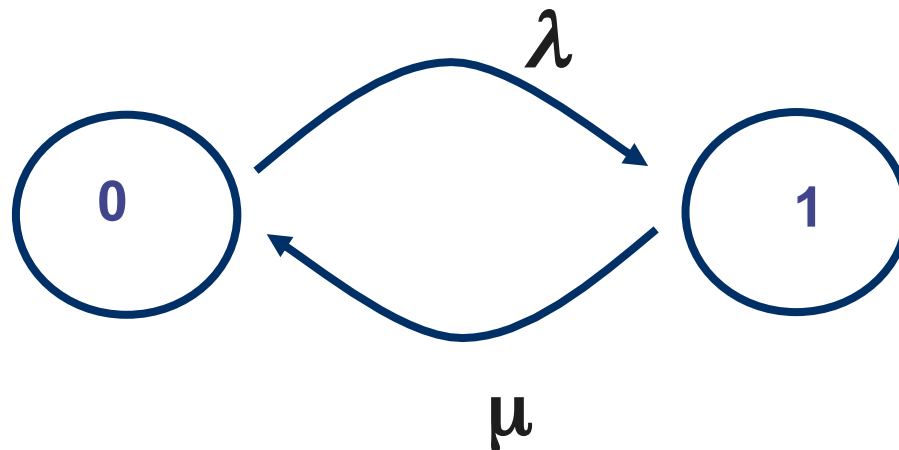
$$\Rightarrow 0 = -\boxed{P_0\lambda}\Delta t + \boxed{P_1\mu}\Delta t$$

Rate of leaving state 0

Rate of entering state 0

Faster way to obtain steady state solution (1)

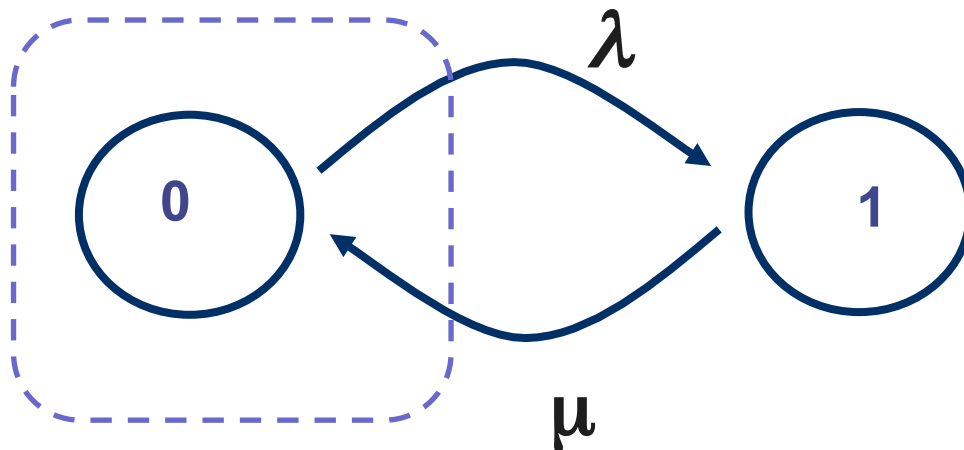
- Transition from State 0 to State 1
 - Caused by an arrival, the rate is λ
- Transition from State 1 to State 0
 - Caused by a completed service, the rate is μ
- State diagram representation
 - *Each circle is a state*
 - *Label the arc between the states with transition rate*



Faster way to obtain steady state solution (2)

- Steady state means
 - **rate of transition out of a state = Rate of transition into a state**
- We have for state 0:

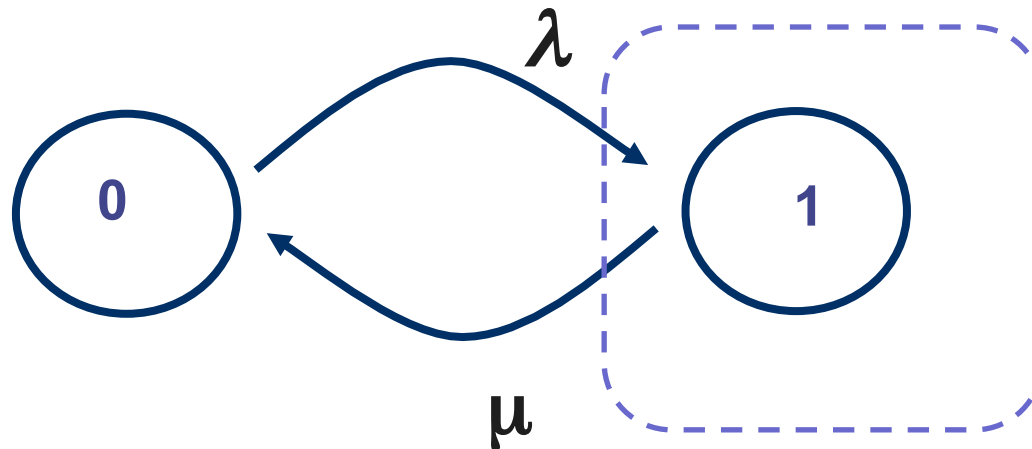
$$\underline{\lambda P_0} = \underline{\mu P_1}$$



Faster way to obtain steady state solution (3)

- We can do the same for State 1:
- Steady state means
 - **Rate of transition into a state** = **rate of transition out of a state**
- We have for state 1:

$$\underline{\lambda P_0} = \underline{\mu P_1}$$



Faster way to obtain steady state solution (4)

- We have one equation $\lambda P_0 = \mu P_1$
- We have 2 unknowns and we need one more equation.
- Since we must be either one of the two states:

$$P_0 + P_1 = 1$$

- Solving these two equations, we get the same steady state solution as before

$$P_0 = \frac{\mu}{\lambda + \mu} \quad P_1 = \frac{\lambda}{\lambda + \mu}$$

Summary

- Solving a queueing problem is not simple
- It is harder to find how a queue evolves with time
- It is simpler to find how a queue behaves at steady state
 - Procedure:
 - Draw a diagram with the states
 - Add arcs between states with transition rates
 - Derive flow balance equation for each state, i.e.
 - Rate of entering a state = Rate of leaving a state
 - Solve the equation for steady state probability

Let us have a look at our call centre problem again

- Consider a call centre
 - Calls are arriving according to Poisson distribution with rate λ
 - The length of each call is exponentially distributed with parameter μ
 - Mean length of a call is $1/\mu$

Call centre:

Arrivals



m operators

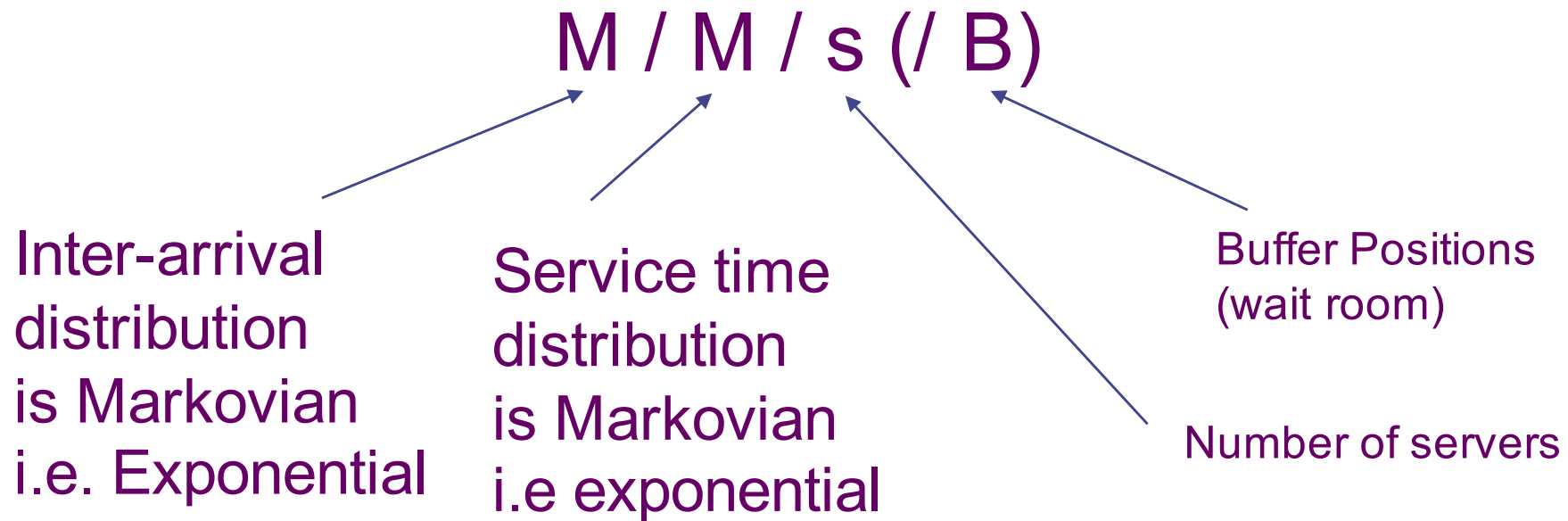
If all operators are busy, the centre can put at most n additional calls on hold.

If a call arrives when all operators and holding slots are used, the call is rejected.

- We solve the problem for $m = 1$ and $n = 0$
 - We call this a M/M/1/1 queue (explanation on the next page)
- How about other values of m and n

Kendall's notation

- To represent different types of queues, queueing theorists use the Kendall's notation
- The call centre example on the previous page can be represented as:



The call centre example on the last page is a $M/M/m/(m+n)$ queue
If $n = \infty$, we simply write $M/M/m$

M/M/1 queue

Exponential
Inter-arrivals (λ)



Exponential
Service time (μ)

Infinite buffer One server

- Consider a call centre analogy
 - Calls are arriving according to Poisson distribution with rate λ
 - The length of each call is exponentially distributed with parameter μ
 - Mean length of a call is $1/\mu$

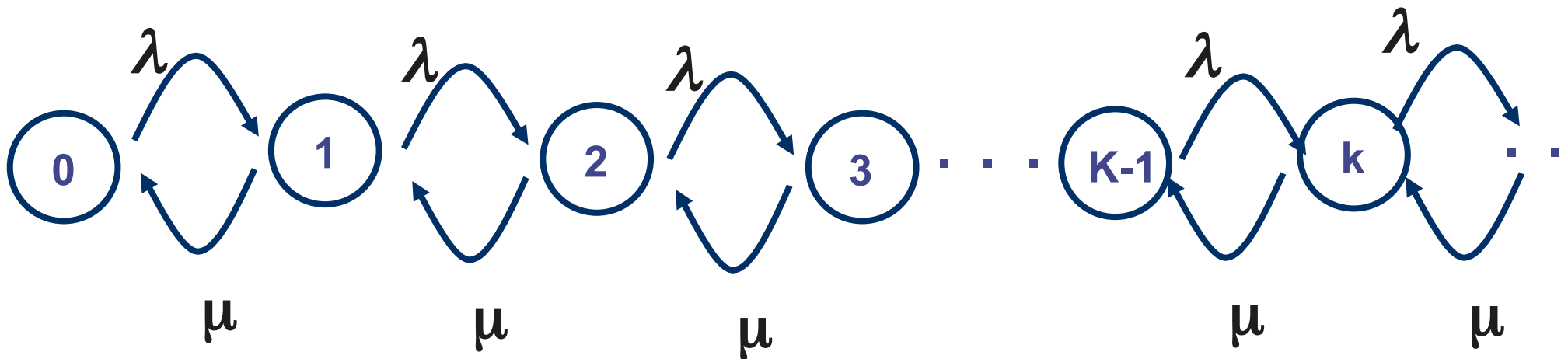
Arrivals
→

Call centre with 1 operator
If the operator is busy, the centre will put the call on hold.
A customer will wait until his call is answered.

- Queueing theory will be able to answer these questions:
 - What is the mean waiting time for a call?

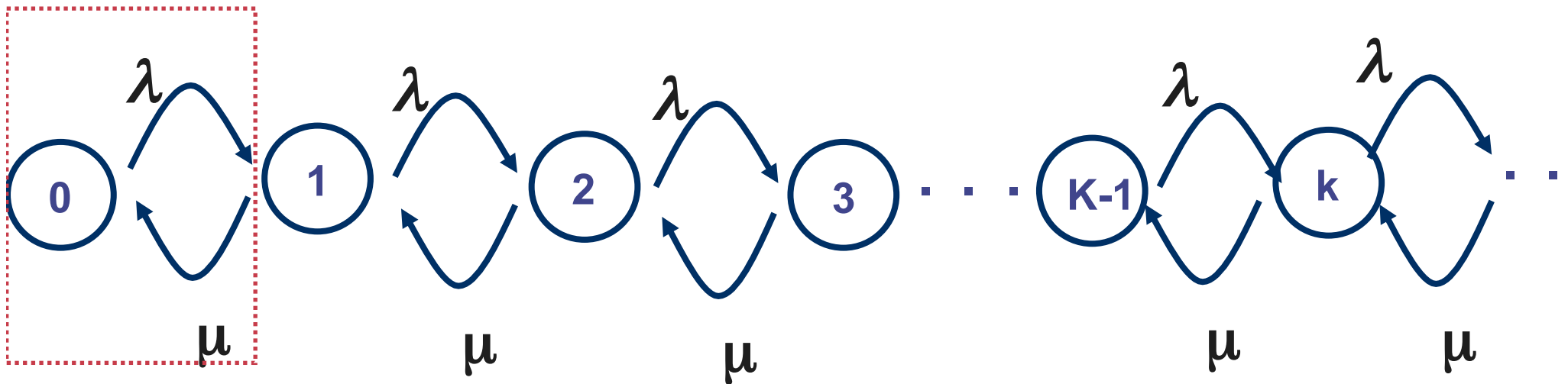
Solving M/M/1 queue (1)

- We will solve for the steady state response
- Define the states of the queue
 - State 0 = There is zero job in the system (= The server is idle)
 - State 1 = There is 1 job in the system (= 1 job at the server, no job queueing)
 - State 2 = There are 2 jobs in the system (= 1 job at the server, 1 job queueing)
 - State k = There are k jobs in the system (= 1 job at the server, $k-1$ job queueing)
- The state transition diagram



Solving M/M/1 queue (2)

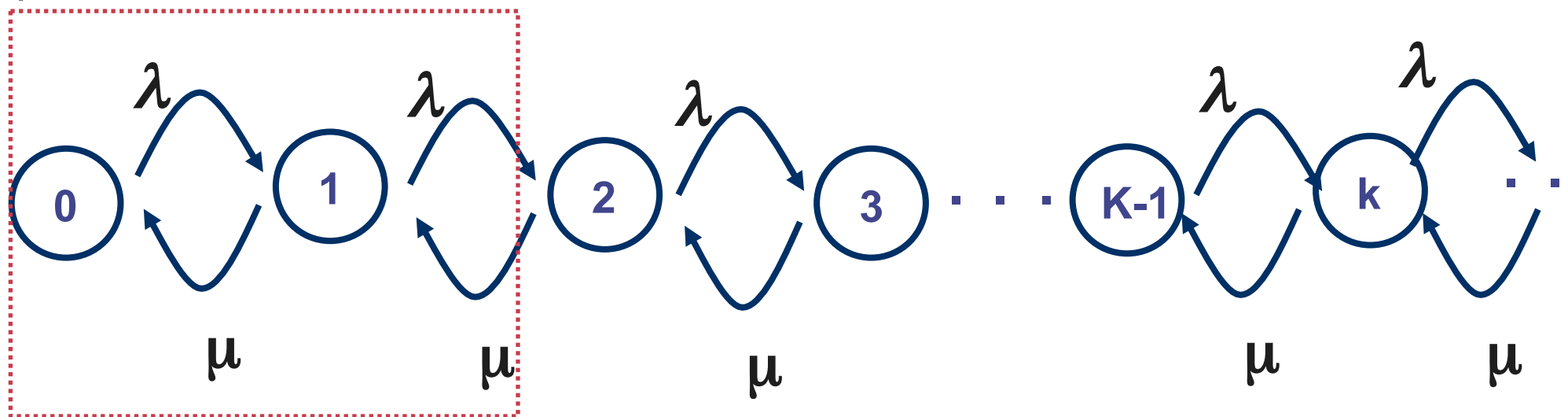
$P_k = \text{Prob. } k \text{ jobs in system}$



$$\lambda P_0 = \mu P_1$$

$$\Rightarrow P_1 = \frac{\lambda}{\mu} P_0$$

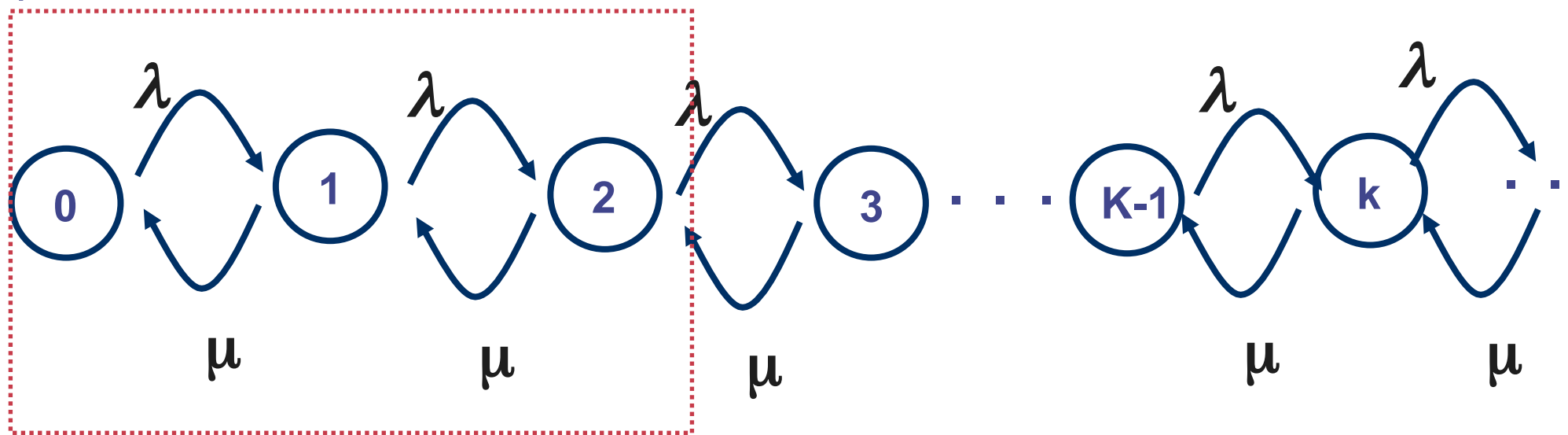
Solving M/M/1 queue (3)



$$\lambda P_1 = \mu P_2$$

$$\Rightarrow P_2 = \frac{\lambda}{\mu} P_1 \quad \Rightarrow P_2 = \left(\frac{\lambda}{\mu} \right)^2 P_0$$

Solving M/M/1 queue (4)



$$\lambda P_2 = \mu P_3$$

$$\Rightarrow P_3 = \frac{\lambda}{\mu} P_2 \quad \Rightarrow P_3 = \left(\frac{\lambda}{\mu} \right)^3 P_0$$

Solving M/M/1 queue (5)

In general $P_k = \left(\frac{\lambda}{\mu}\right)^k P_0$

Let $\rho = \frac{\lambda}{\mu}$

We have $P_k = \rho^k P_0$

Solving M/M/1 queue (6)

With $P_k = \rho^k P_0$ and

$$P_0 + P_1 + P_2 + P_3 + \dots = 1$$

$$\Rightarrow (1 + \rho + \rho^2 + \dots)P_0 = 1$$

$$\Rightarrow P_0 = 1 - \rho \text{ if } \rho < 1$$

$$\Rightarrow P_k = (1 - \rho)\rho^k$$

Since $\rho = \frac{\lambda}{\mu}$, $\rho < 1 \Rightarrow \lambda < \mu$

ρ = utilisation
= Prob server is busy
= $1 - P_0$
= 1- Prob server is idle

Arrival rate < service rate

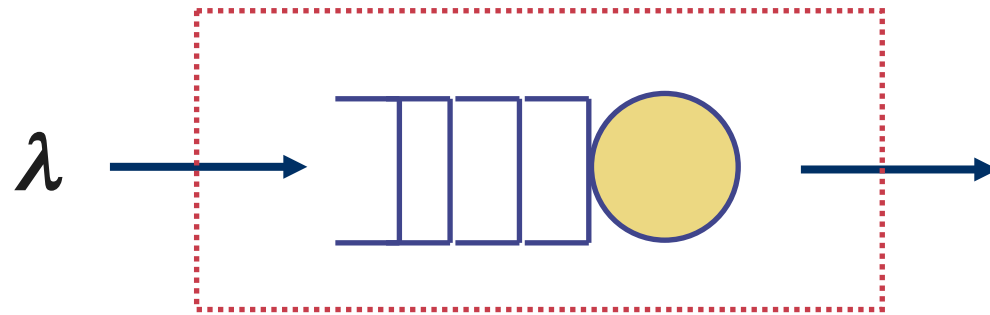
Solving M/M/1 queue (7)

With $P_k = (1 - \rho)\rho^k$

This is the probability that there are k jobs in the system.
To find the response time, we will make use of Little's law.
First we need to find the mean number of customers =

$$\begin{aligned}\sum_{k=0}^{\infty} k P_k &= \sum_{k=0}^{\infty} k (1 - \rho) \rho^k \\ &= \frac{\rho}{1 - \rho}\end{aligned}$$

Solving M/M/1 queue (8)



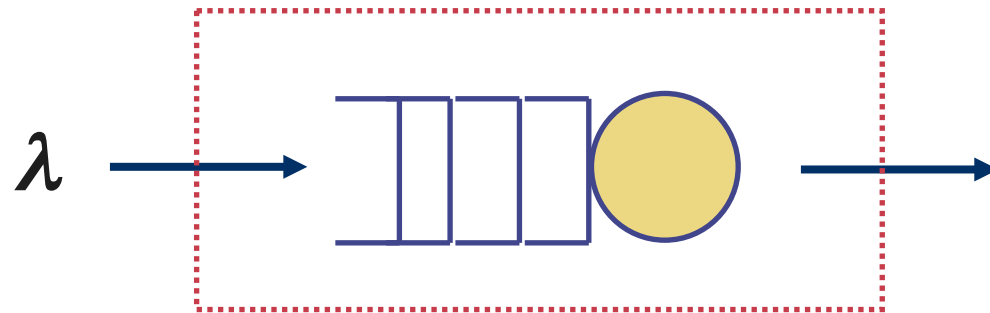
Little's law:

mean number of customers = throughput x response time

Throughput is λ (*why?*)

$$\text{Response time } T = \frac{\rho}{\lambda(1 - \rho)} = \frac{1}{\mu - \lambda}$$

Solving M/M/1 queue (9)



What is the mean waiting time at the queue?

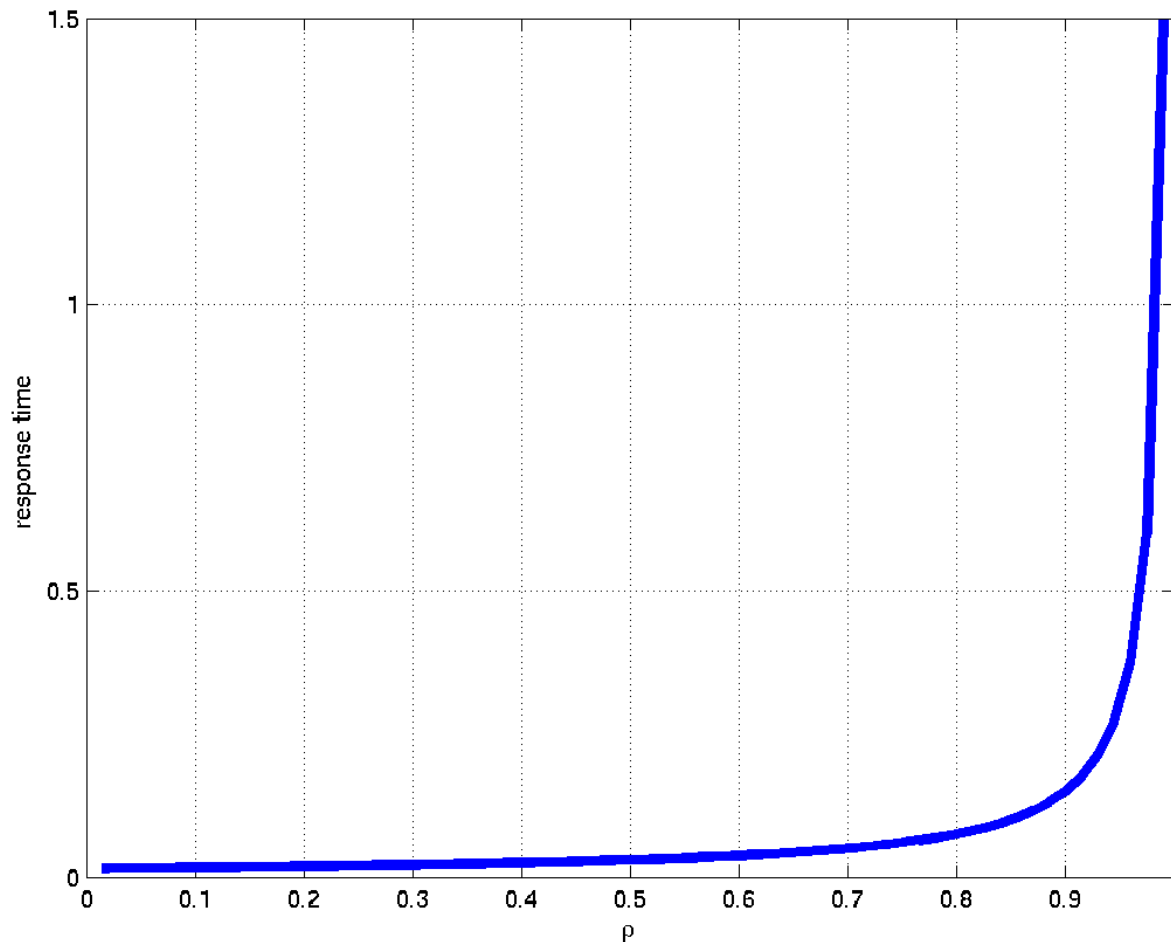
Mean waiting time = mean response time - mean service time

We know mean response time (from last slide)

Mean service time is = $1 / \mu$

Using the service time parameter ($1/\mu = 15\text{ms}$) in the example, let us see how response time T varies with λ

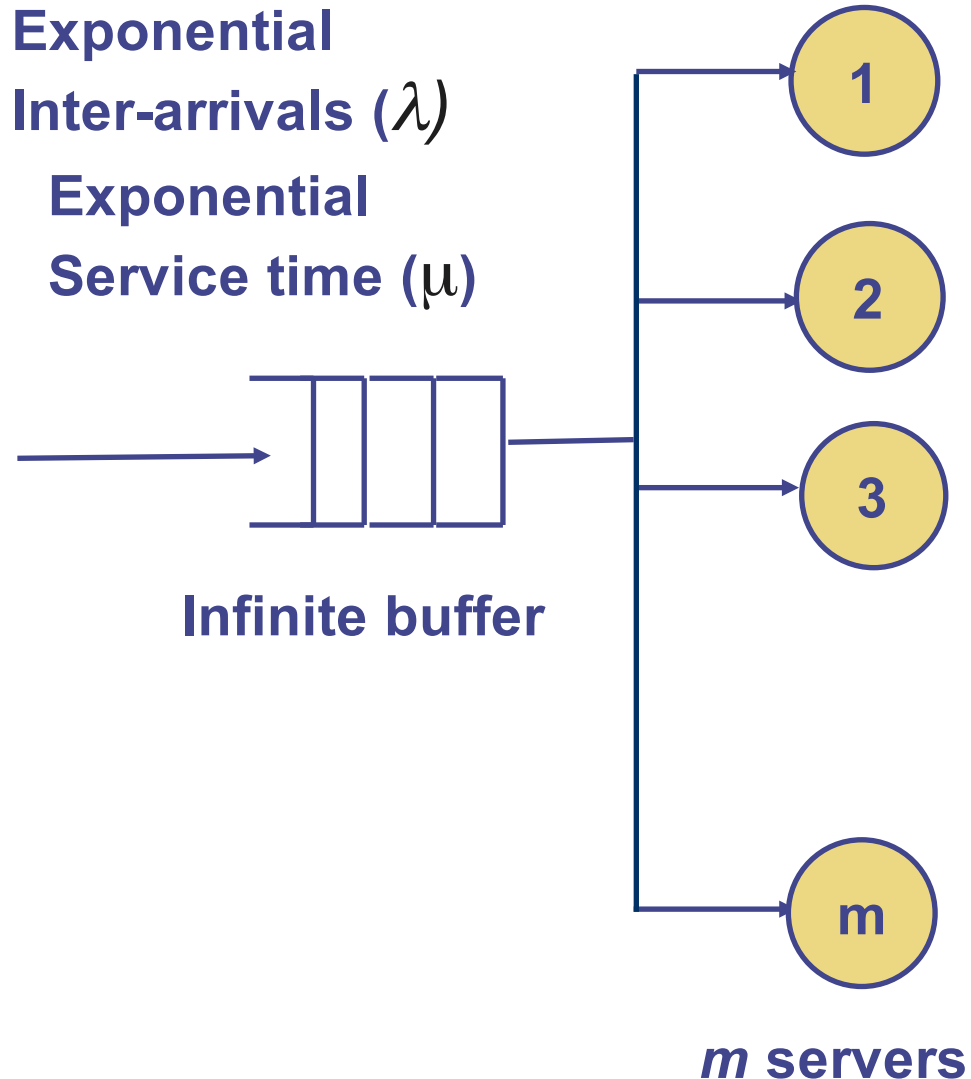
$$T = \frac{1}{\mu(1 - \rho)}$$



Observation:
Response time increases sharply when ρ gets close to 1

Infinite queue assumption means $\rho \rightarrow 1$, $T \rightarrow \infty$

Multi-server queues M/M/m



All arrivals go into one queue.

Customers can be served by any one of the m servers.

When a customer arrives

- If all servers are busy, it will join the queue
- Otherwise, it will be served by one of the available servers

A call centre analogy of M/M/m queue

- Consider a call centre analogy
 - Calls are arriving according to Poisson distribution with rate λ
 - The length of each call is exponentially distributed with parameter μ
 - Mean length of a call is $1/\mu$

Arrivals

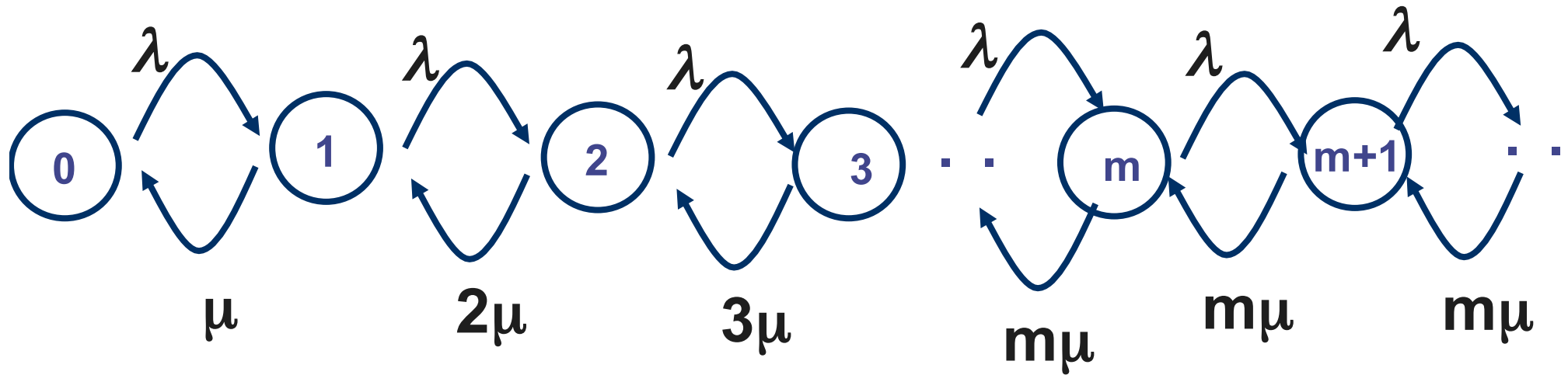


Call centre with m operators

If all m operators are busy, the centre will put the call on hold.

A customer will wait until his call is answered.

State transition for M/M/m



M/M/m

- Following the same method, we have mean response time T is

$$T = \frac{C(\rho, m)}{m\mu(1 - \rho)} + \frac{1}{\mu}$$

where

$$\rho = \frac{\lambda}{m\mu}$$

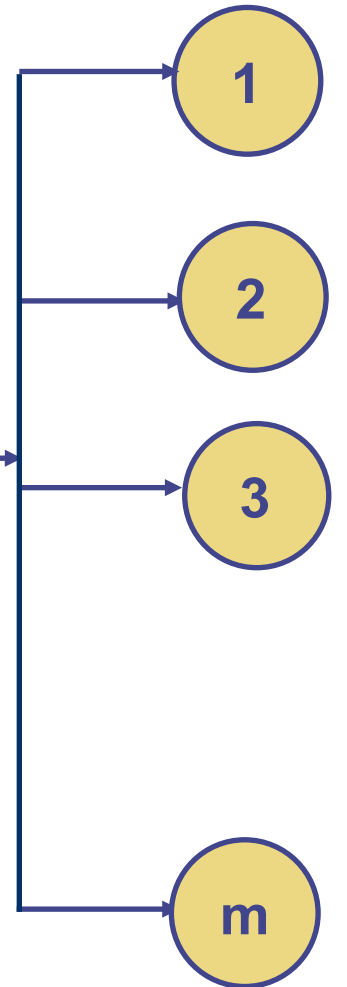
$$C(\rho, m) = \frac{\frac{(m\rho)^m}{m!}}{(1 - \rho) \sum_{k=0}^{m-1} \frac{(m\rho)^k}{k!} + \frac{(m\rho)^m}{m!}}$$

Multi-server queues M/M/m/m with no waiting room

Exponential
Inter-arrivals (λ)

Exponential
Service time (μ)

No waiting
Room or
No buffer



***m* servers**

An arrival can be served by any one of the m servers.

When a customer arrives

- If all servers are busy, it will *depart* from the system

- Otherwise, it will be served by one of the available servers

A call centre analogy of M/M/m/m queue

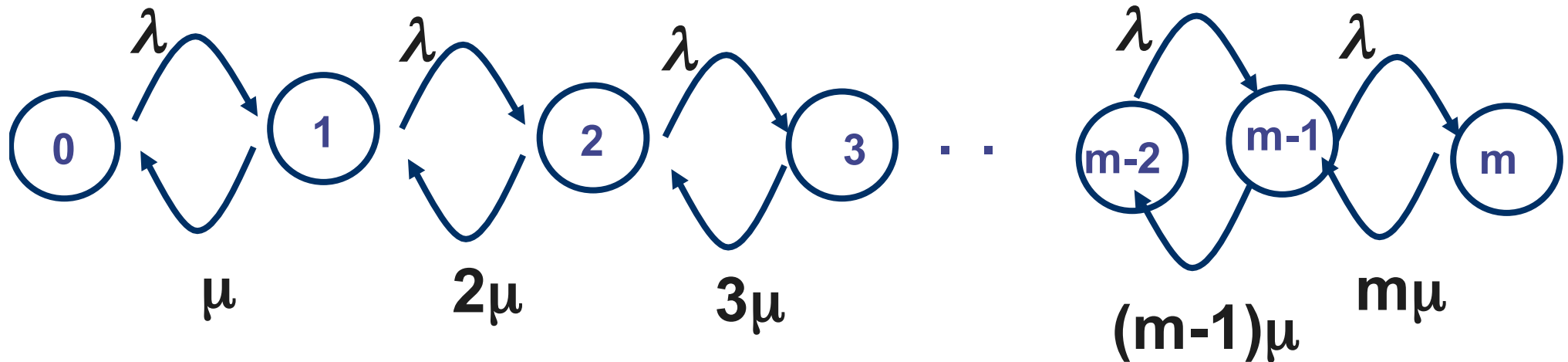
- Consider a call centre analogy
 - Calls are arriving according to Poisson distribution with rate λ
 - The length of each call is exponentially distributed with parameter μ
 - Mean length of a call is $1/\mu$

Arrivals



Call centre with m operators
If all m operators are busy, the call is dropped.

State transition for M/M/m/m



**Probability that an arrival is blocked
= Probability that there are m customers in the system**

$$P_m = \frac{\frac{\rho^m}{m!}}{\sum_{k=0}^m \frac{\rho^k}{k!}} \quad \text{where} \quad \rho = \frac{\lambda}{\mu}$$

“Erlang B formula”

Poisson arrivals see time averages (PASTA)

- P_n = Probability that there are n jobs in the system
- A_n = Probability that an arriving customer finds n jobs in the system
- If the arrival process is Poisson, then $A_n = P_n$
- Proof: Need to show the following two expressions are equal.

$$A_n = \lim_{t \rightarrow \infty} \lim_{\delta \rightarrow 0} \text{Prob} [n \text{ jobs in the system at time } t \mid \text{an arrival occurs in } (t, t + \delta)]$$

$$P_n = \lim_{t \rightarrow \infty} \text{Prob} [n \text{ jobs in the system at time } t]$$

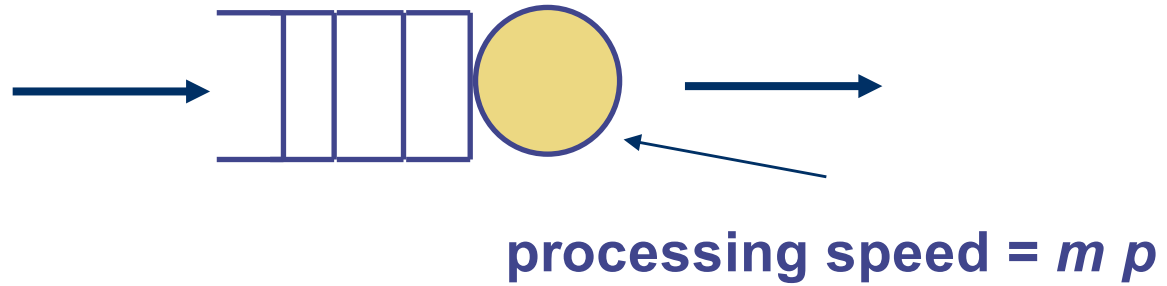
- Key step in the proof, Poisson arrival means

$$\begin{aligned} & \text{Prob} [\text{an arrival occurs in } (t, t + \delta) \mid n \text{ jobs in the system at time } t] \\ &= \text{Prob} [\text{an arrival occurs in } (t, t + \delta)] \end{aligned}$$

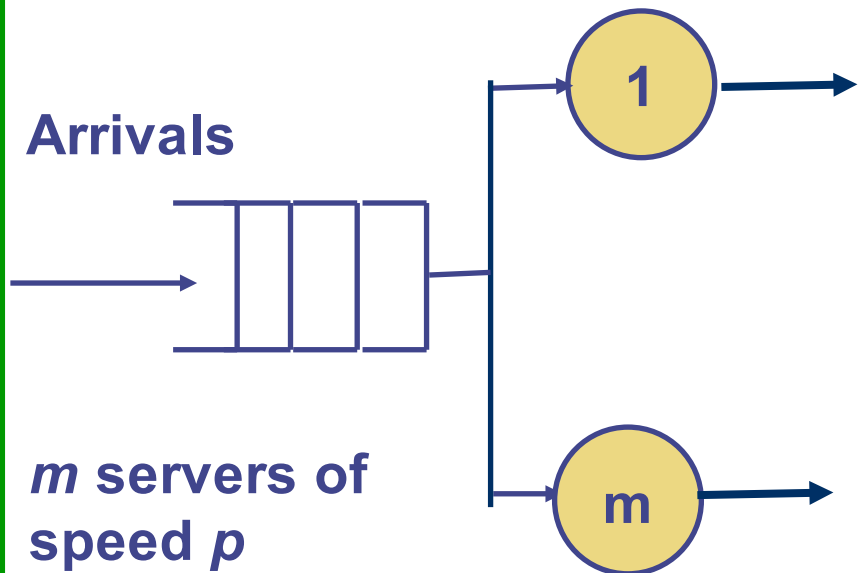
- To be completed in class

What configuration has the best response time?

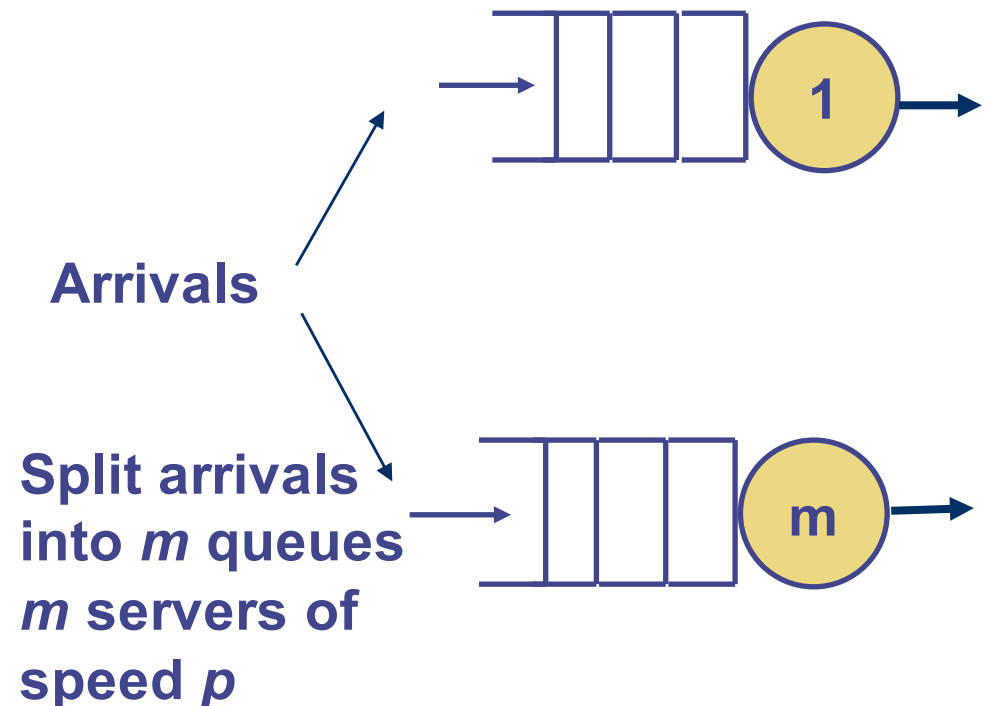
Configuration 1:



Configuration 2:



Configuration 3:



Try out the tutorial question!

References

- Recommended reading
 - Queues with Poisson arrival are discussed in
 - Bertsekas and Gallager, *Data Networks*, Sections 3.3 to 3.4.3
 - Note: I derived the formulas here using continuous Markov chain but Bertsekas and Gallager used discrete Markov chain
 - Mor Harchal-Balter. Chapters 13 and 14
 - Poisson arrival sees time averages (PASTA)
 - See R.W. Wolff, “Poisson Arrivals See Time Averages”, *Operational Research*, Vol 30, No 2, pp.223-231
 - (Accessible within UNSW) www.jstor.org/stable/170165